

# Gene Expression in Fourth and Fifth Instar *Manduca Sexta*

Robert A. Long and Bradley A. Hartlaub, Department of Mathematics, Kenyon College, Gambier, OH 43022

## Abstract

- The scaling exponent for metabolic rate in *Manduca sexta* is greater than expected based on isometric surface area scaling. We examined whether midgut genes involved in absorption and digestion are expressed differently in fifth versus fourth instar larvae. RNA was isolated from middle and posterior midgut and reverse transcribed to cDNA. Real-time PCR was used to quantify expression of two genes, aminopeptidase N (*msAPN3*) and potassium amino acid cotransporter (*KAAT*), by the relative quantification method using 18s ribosomal RNA as an internal control. A general linear model for  $\Delta Ct$  values from *APN* and *KAAT* identified significant instar and tissue effects.
- The best model for *KAAT1* and *KAAT2* provided a good fit, with an R-squared value of 0.888431. In posterior midgut, *KAAT* expression was 2.5-3.5 fold higher while *APN* was 1.3-1.4 fold lower in fifth compared to fourth instar larvae. Expression of *KAAT* was 3.5 fold higher in middle compared to posterior midgut.
- To evaluate the sensitivity of standard statistical methods to real time PCR data, a Monte Carlo simulation of the coverage probabilities for Ct,  $\Delta Ct$  and  $\Delta\Delta Ct$  was conducted for the normal, uniform, exponential, t and chi-square distributions using parameter estimates derived from our data. Coverage probabilities for Ct and  $\Delta Ct$  were close to the 95% confidence level, indicating that the interval estimates are relatively insensitive to distributional assumptions. A notable exception was the exponential distribution of Ct, where the coverage probability dropped to 0.88. Coverage probabilities for  $\Delta\Delta Ct$  were reasonable when using the mean of a comparison group, but not when arbitrarily selecting an individual measurement for comparison.

## Data Collection

### Data Collection:

- The data can be subdivided by: Cohort (a group of *Manduca Sexta* grown and analyzed together), Plate, Primer (*APN*, *KAAT* or *18s*), Tissue (MM or PM), and Instar (4<sup>th</sup> or 5<sup>th</sup>).
- For each sample, six Ct values are measured (three with *18s* and three with *APN* or *KAAT*).
  - Replicates of three are a necessity in order to have a reasonable power (see Figure 1).
- Ct is the number of cycles necessary for the amount of gene expressed to meet a predetermined threshold (see Figure 2).
- Using these Ct values, we can compute  $\Delta Ct$  and  $\Delta\Delta Ct$  where  $\Delta Ct = Ct_{KAAT} - Ct_{18s}$  and  $\Delta\Delta Ct = \Delta Ct - \Delta Ct^*$ .
  - For  $\Delta\Delta Ct$  there does not exist a clear treatment-control relationship within the data. Computer software will tend to use an arbitrary observation as the control ( $\Delta Ct^*$ ); we have found that there are problems with this technique.

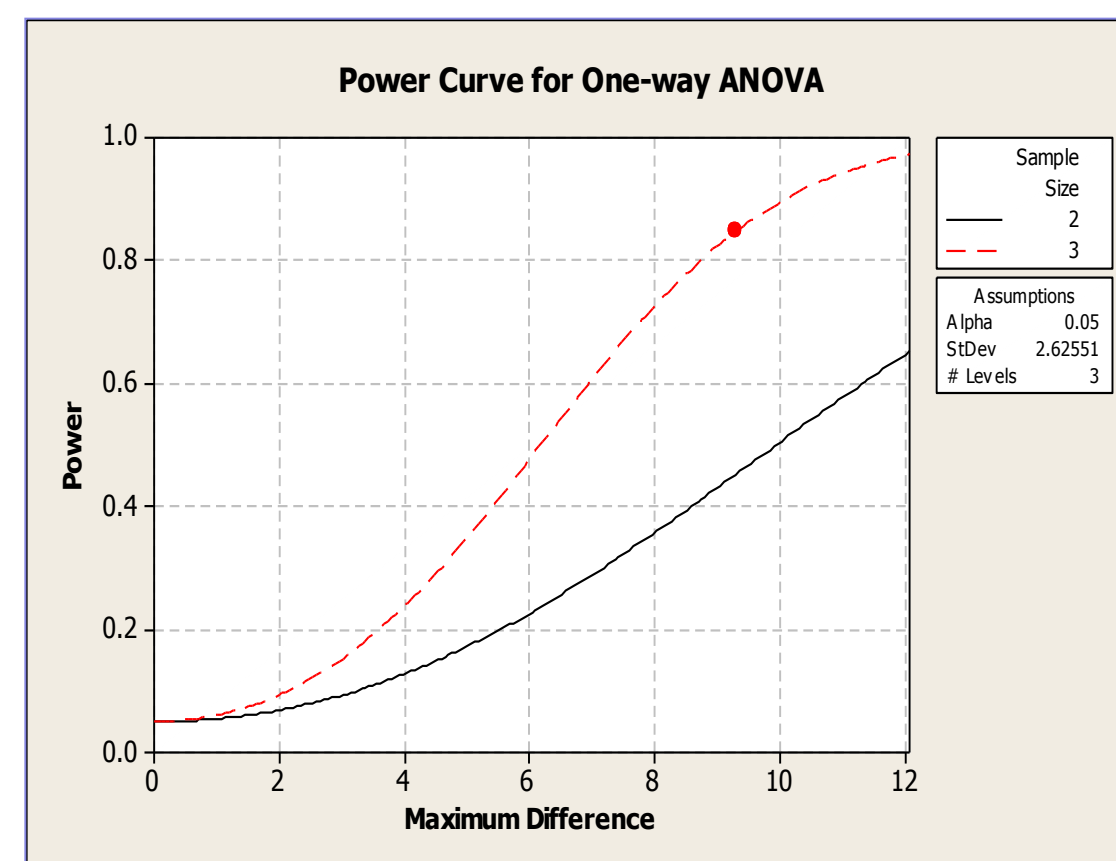


Figure 1: Power Curve for Replicates

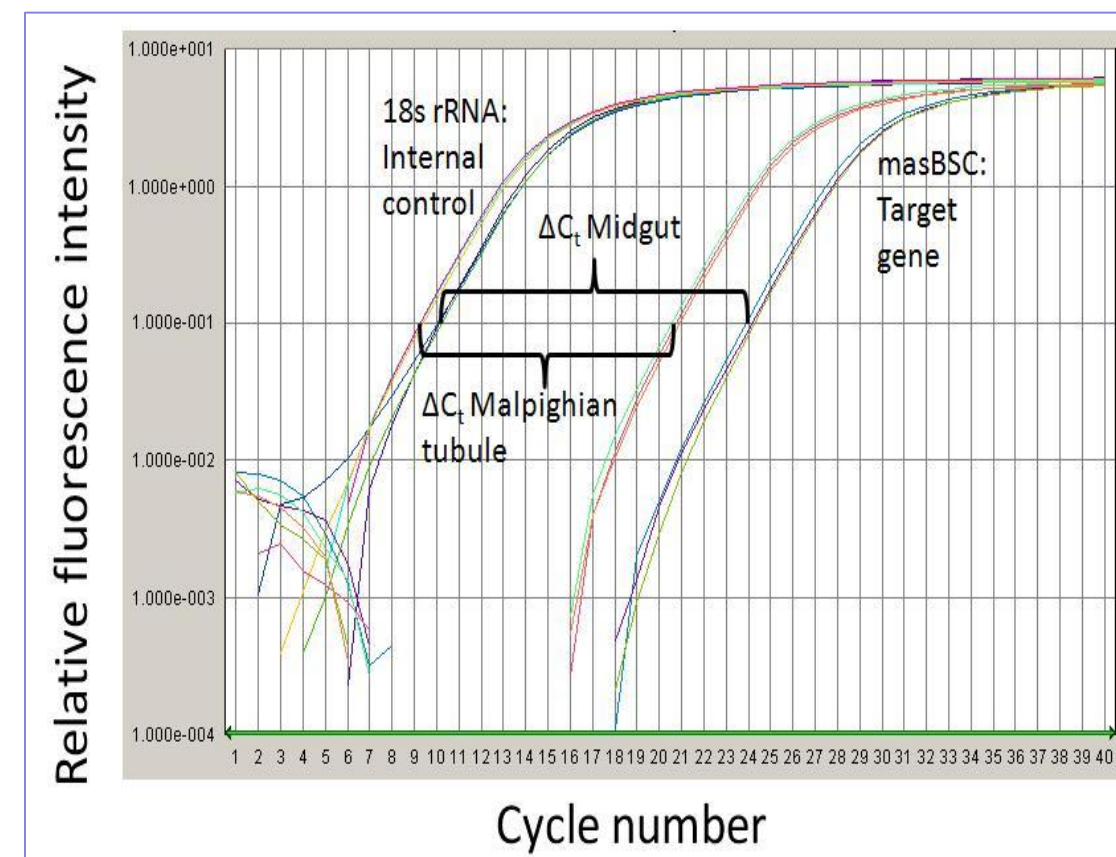


Figure 2: Ct and  $\Delta Ct$  Expression

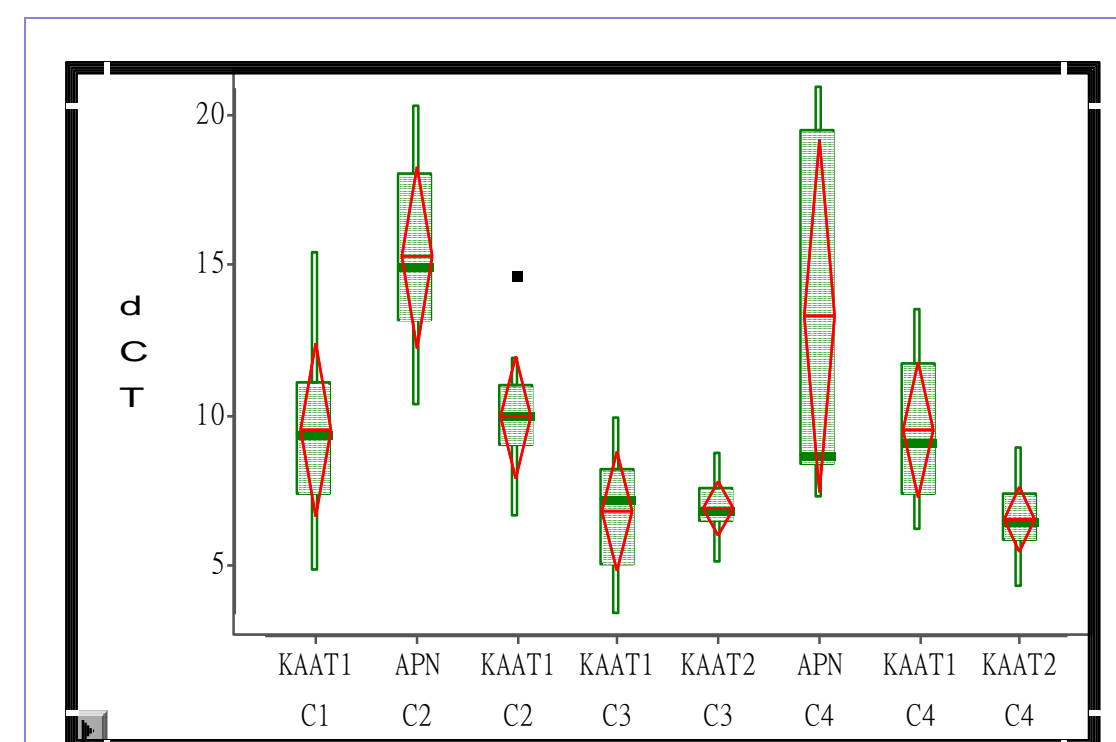


Figure 3:  $\Delta Ct$  vs. Cohort, Primer

## Statistical Inference Results

- Anderson-Darling:  $H_0$ : Data ( $\Delta Ct$ ) are normal,  $H_a$ : Data are not Normal
  - Test statistic: A-sq = 7.5525, P-value: 0.0050
  - However, when the data are separated by primer, tissue and cohort many of the subgroups confirm normality.
- T-Tests show significant differences between most cohorts (see Figures 3, 8).
- $H_0: \mu_i = \mu_j$  vs.  $H_a: \mu_i \neq \mu_j$  (where  $\mu_i$  is the mean  $\Delta Ct$  for cohort  $i$ )
- The F-test for equal variance shows significant differences between primers.
  - $H_0: \sigma_i^2 = \sigma_j^2$ ,  $H_a: \sigma_i^2 \neq \sigma_j^2$  where  $\sigma_i^2$  is the variance of  $\Delta Ct$  for primer  $i$ .

Cohort	N	Mean	SD	Min	Median	Max
1	23	9.48	2.87	4.79	9.34	15.37
2	32	12.59	3.71	6.54	11.70	20.22
3	54	6.82	1.65	3.38	6.88	9.83
4	73	9.87	4.36	4.42	8.36	20.96

Primer	N	Mean	SD	Min	Median	Max
KAAT1	104	8.70	2.62	3.38	8.38	15.37
KAAT2	40	6.72	0.99	4.24	6.62	8.84
APN	38	14.12	4.90	7.25	14.17	20.96

Primer <sub>i</sub>	Primer <sub>j</sub>	Num. DF	Den. DF	F-Value	P-Value
KAAT1	KAAT2	103	39	7.03	<.0001
APN	KAAT1	37	103	3.50	<.0001
APN	KAAT2	37	39	24.58	<.0001

## General Linear Model

- $\Delta Ct_i = \beta_0 + \beta_1 * Plate_i + \beta_2 * Tissue_i + \beta_3 * Primer_i + \beta_4 * Instar_i + \beta_5 * Instar_i * Plate_i + \epsilon_i$
- Instar and the interaction between plate and instar were found to be significant (see figure 9). As expected, Plate, Tissue (MM, PM) and Primer (*KAAT1*, *KAAT2*) were also found to be significant.
- $R^2 = 0.888431$

Source	DF	Type III SS	Mean Square	F Statistic	P-Value
Plate	17	489.888	28.81694	31.88	<.0001
Tissue	1	21.454	21.48505	23.77	<.0001
Primer	1	15.065	15.06456	16.66	<.0001
Instar	1	20.018	20.01758	22.14	<.0001
Instar*Plate	17	95.289	5.60523	6.20	<.0001

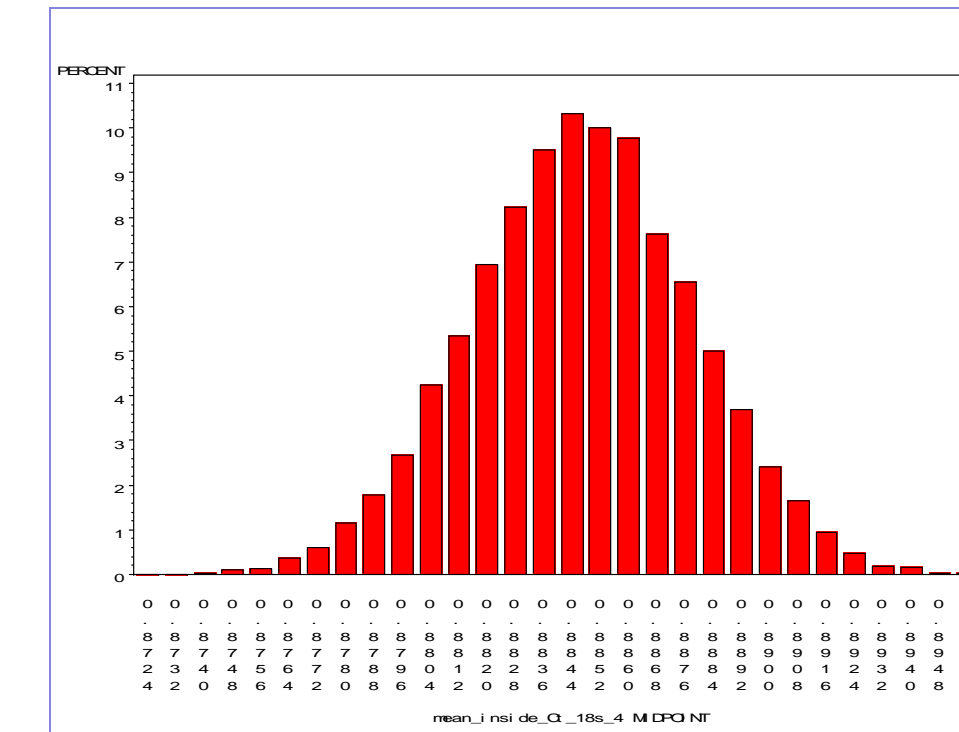


Figure 4: Ct for Exponential

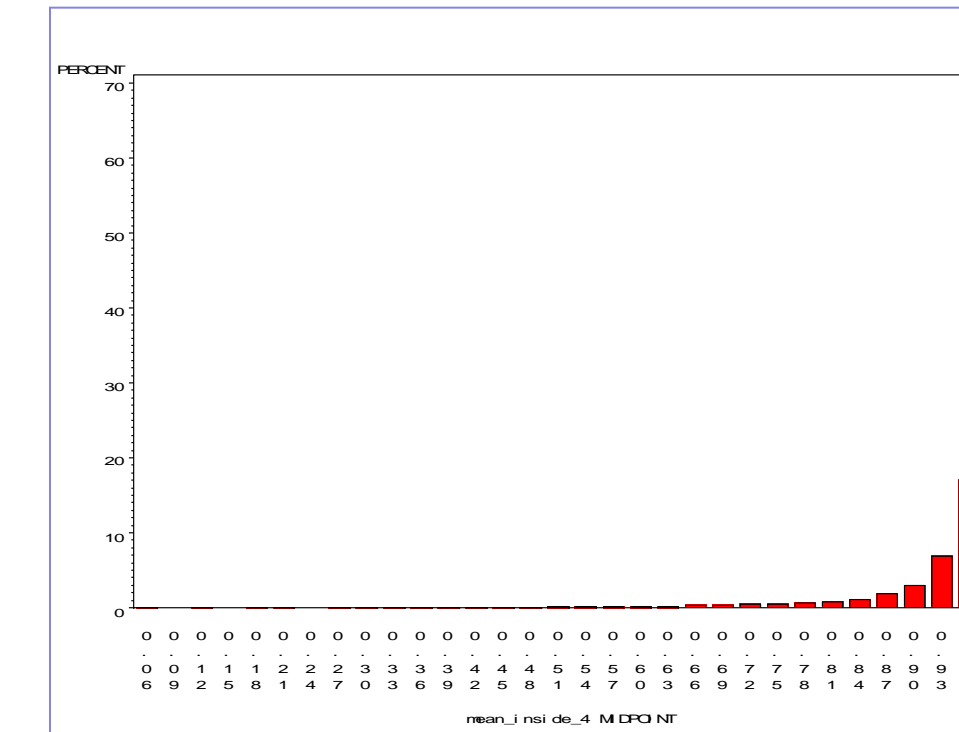


Figure 5:  $\Delta\Delta Ct$  for Exponential with Random

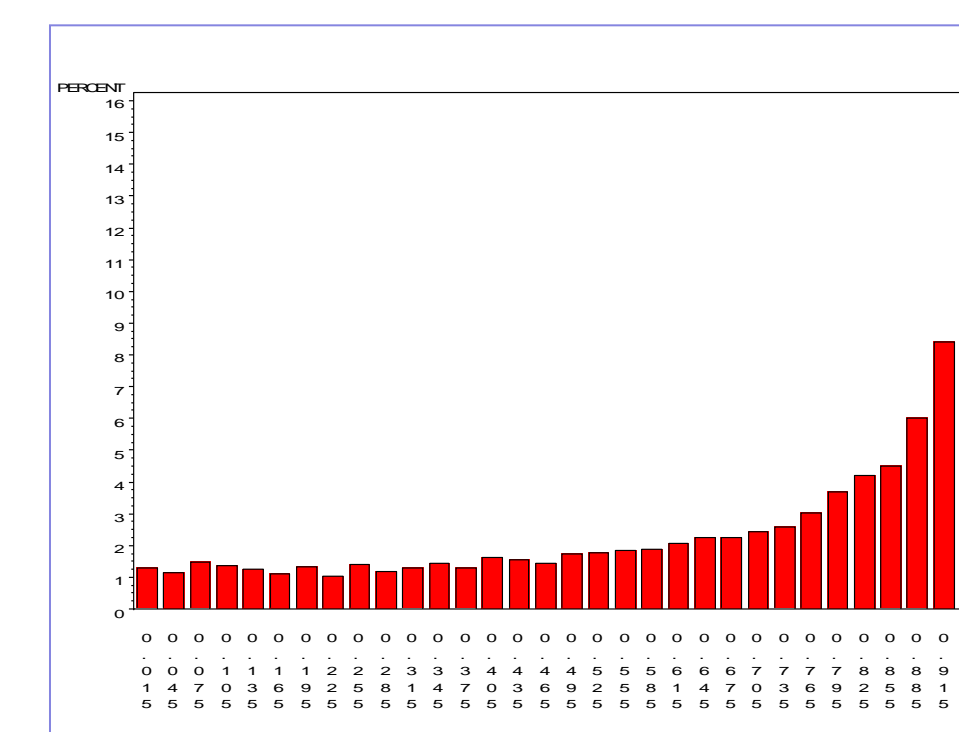


Figure 6:  $\Delta\Delta Ct$  for t with Mean

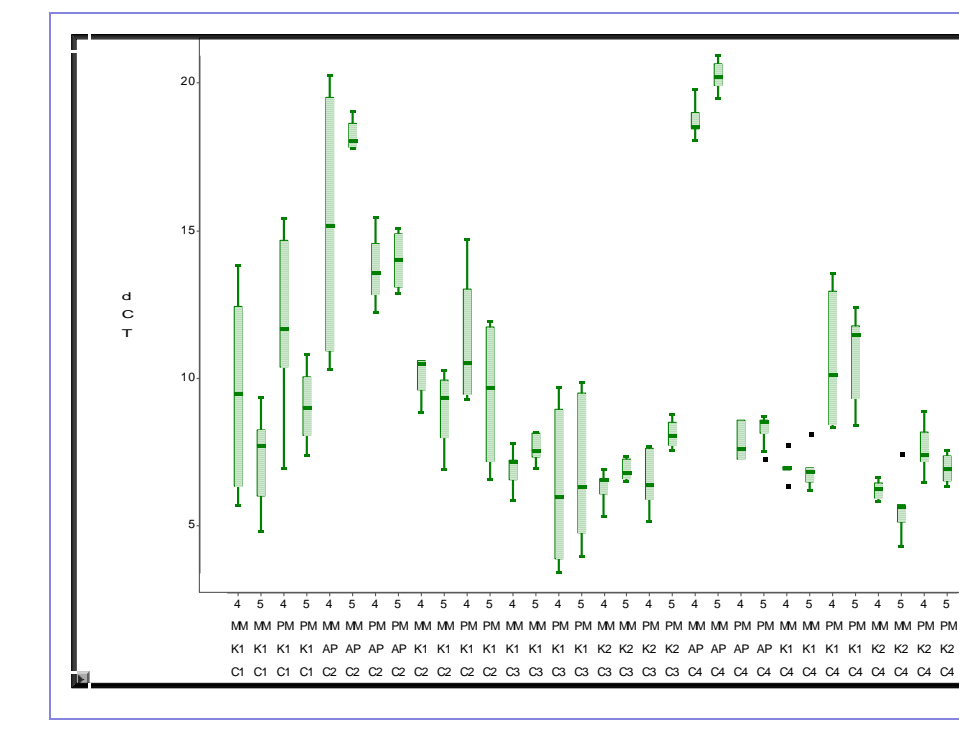


Figure 7:  $\Delta Ct$  vs. Cohort, Primer, Tissue and Instar

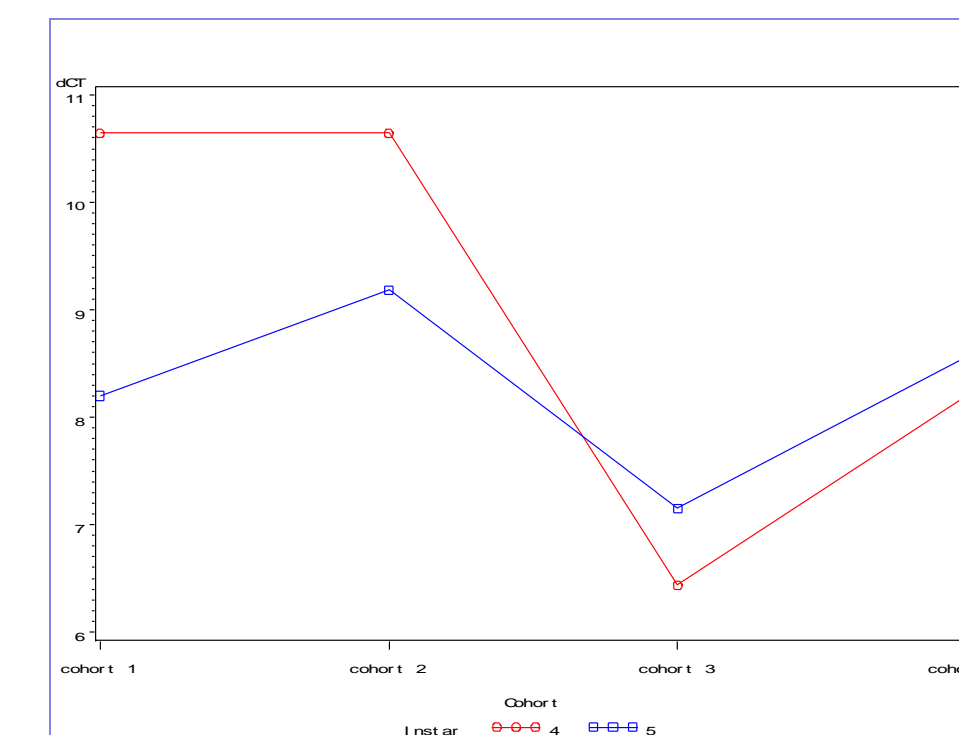


Figure 8: Profile Plot of  $\Delta Ct$  vs Cohort (*KAAT1*, *KAAT2*)

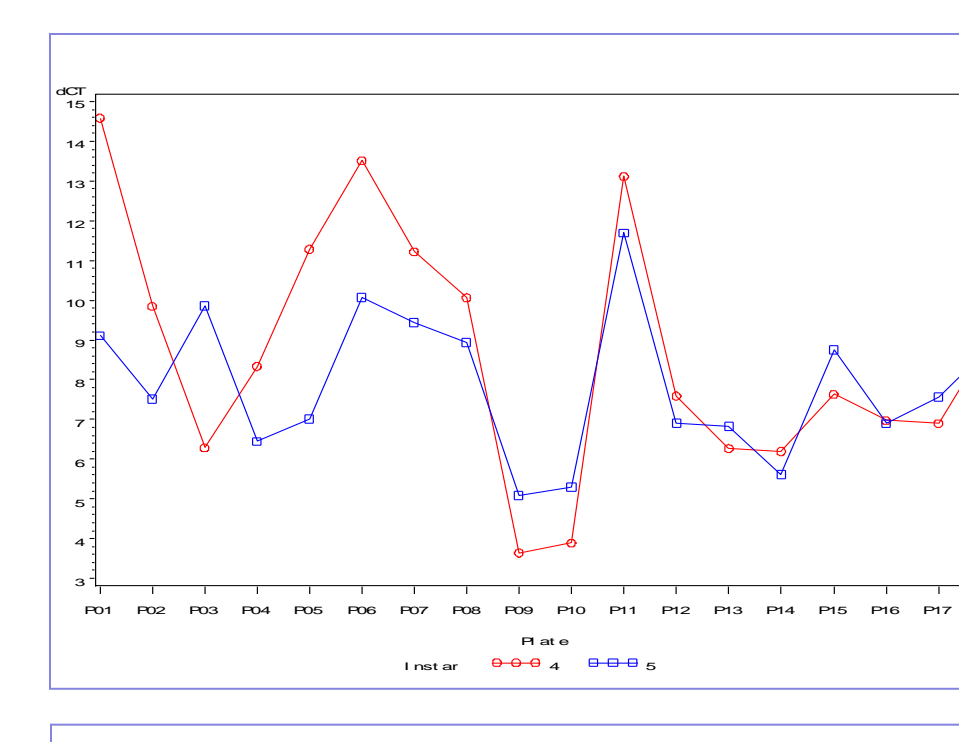


Figure 9: Profile Plot of  $\Delta Ct$  vs Plate by Instar (*KAAT1*, *KAAT2*)

## Monte Carlo Simulation

### Methods:

- Coverage probabilities were simulated for Ct,  $\Delta Ct$  and  $\Delta\Delta Ct$  for 4<sup>th</sup> and 5<sup>th</sup> instars.
- Ct values were randomly generated as either Normal, Uniform, Exponential, t or Chi-square using parameter estimates derived from our data.
- We then constructed a confidence interval and checked whether the mean was within the bounds.
  - This procedure was repeated 10,000 times to estimate the coverage probability.
  - Then 10,000 coverage probabilities were calculated and we created a histogram to show the distribution of the coverage probabilities.

### Results:

- Ct was found to be relatively insensitive to changes in distribution.
  - Notable exceptions occurred with the uniform (coverage dropped to around 0.92) and the exponential (around 0.885)
- $\Delta Ct$  saw a general increase in coverage (normal = 0.957)
  - Exponential improved to 0.955 and uniform to 0.954
- Three different controls were used for computing  $\Delta\Delta Ct$ :
  - An arbitrary  $\Delta Ct$  value
    - The coverage, though often centered around 0.95, had an extremely large range (from 0.04 to 1).
  - The mean of the 4<sup>th</sup> instar  $\Delta Ct$  values
    - This method showed excellent coverage, though slightly below 0.95, the range was small and the coverage estimates were normally distributed. This method was unsuccessful for the t-distribution because of the extreme outliers.
  - The median of the 4<sup>th</sup> instar  $\Delta Ct$  values
    - Unaffected by the outliers in the t-distribution.

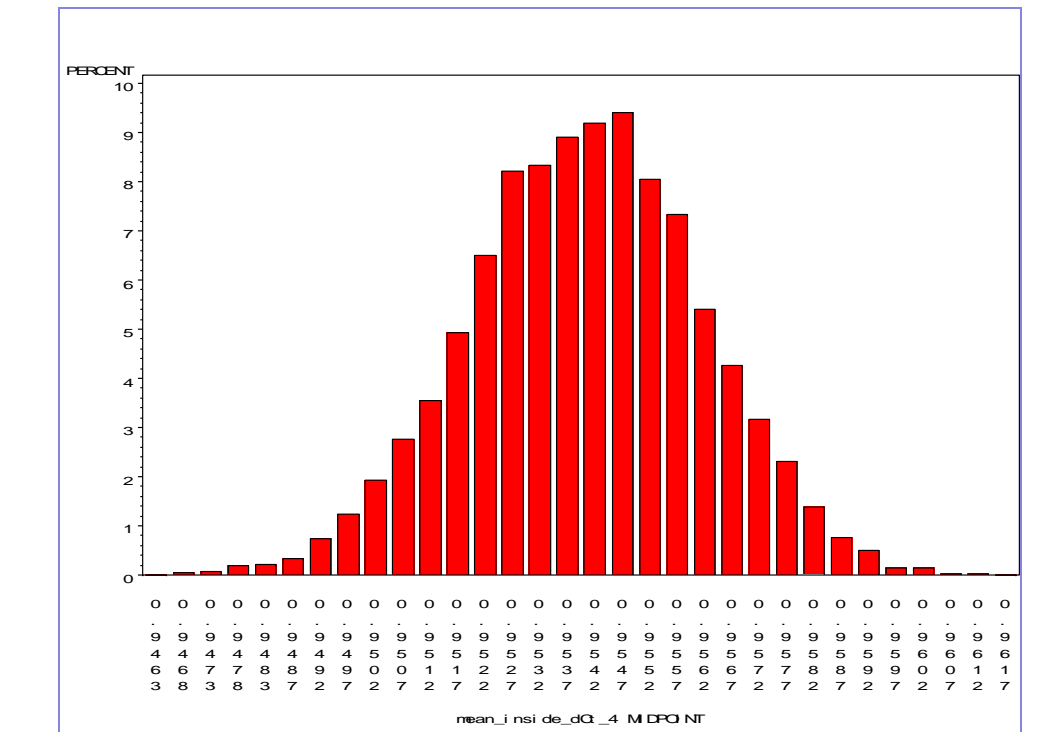


Figure 10:  $\Delta Ct$  for Exponential

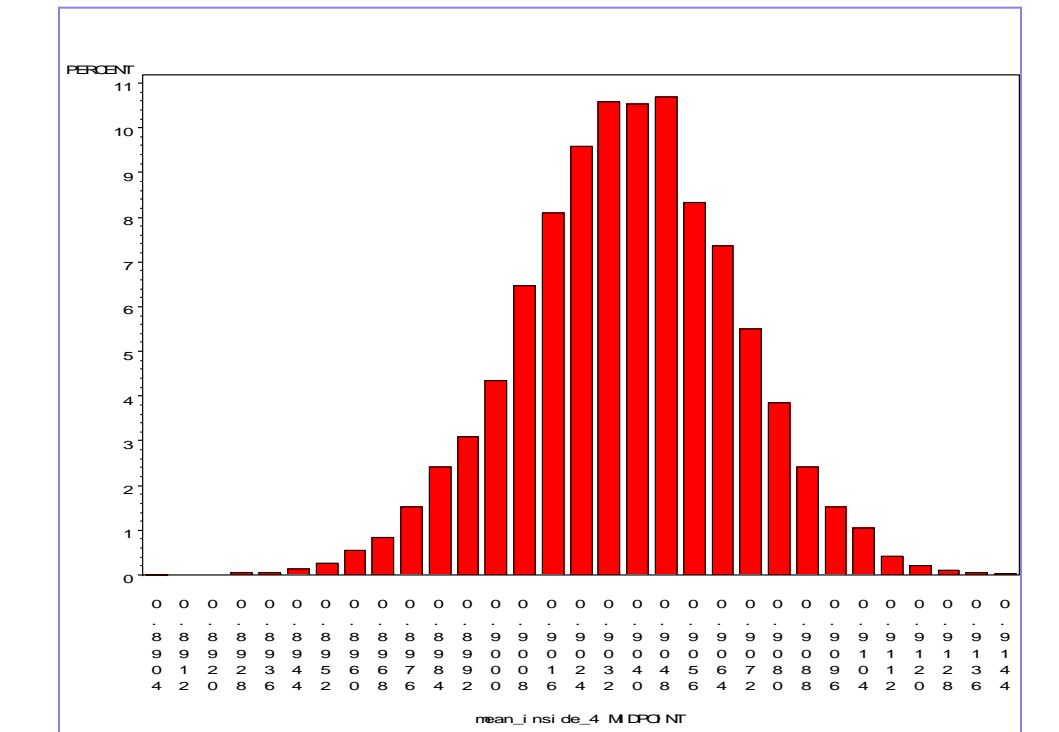


Figure 11:  $\Delta\Delta Ct$  for Exponential with Mean

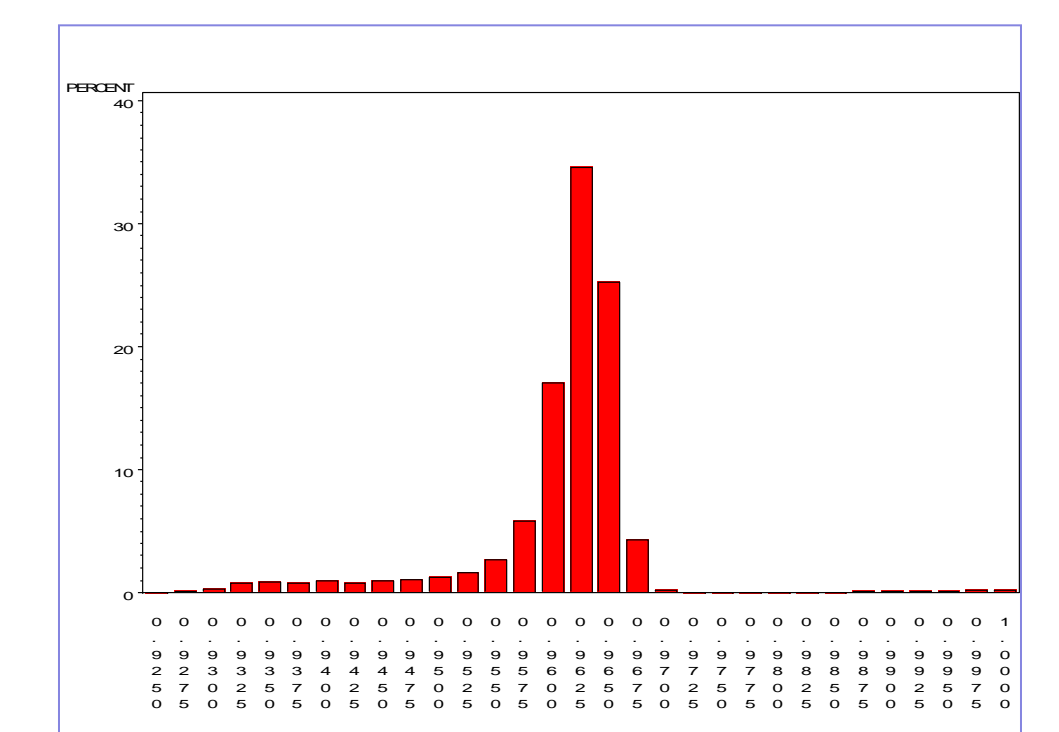


Figure 12:  $\Delta\Delta Ct$  for t with Median

## Conclusions

- The general linear model showed significant differences in the expression of middle and posterior midgut genes in 4<sup>th</sup> and 5<sup>th</sup> instar *Manduca Sexta*.
- As expected there were also differences in the expression based on tissue, plate and primer.
- The Monte Carlo simulation confirmed  $\Delta Ct$  as a more reliable measurement than Ct.
  - Ct and  $\Delta Ct$  were both relatively insensitive to distributional changes though key exceptions were in the uniform and exponential distribution
  - Of the three ways to compute  $\Delta\Delta Ct$ , choosing a random observation as the control was shown to be unreliable.
  - The mean worked well as long as we were not dealing with the t-distribution where we needed to use the median.
  - Future research should consider nonparametric alternatives because of the non-normality of our data. The Monte Carlo simulation could also be extended to simulate fold changes (defined as  $2^{-\Delta\Delta Ct}$ ).

## References

- Yuan, Joshua S., Ann Reed, Feng Chen and C. Neal Stewart Jr. "Statistical Analysis of Real-Time PCR Data." *BMC Bioinformatics* 7:85 (22 February 2006):
- Gibson, Ursula E.M., Christian A. Heid, and P. Mickey Williams. "A Novel Method for Real Time Quantitative RT-PCR." *Genome Research* 6 (16 August 1996): 995-1001.
- Karlen, Yann, Alan McNair, Sebastien Perseguers, Christian Mazza and Nicolas Mermod. "Statistical Significance of Quantitative PCR." *BMC Bioinformatics* 8:131 (20 April 2007).
- Livak, Kenneth J. and Thomas D. Schmittgen. "Analysis of Relative Gene Expression Data Using Real-Time Quantitative PCR and the  $2^{-\Delta\Delta Ct}$  Method." *Methods* 25 (2001): 402-408.

## Acknowledgements

I would like to thank Brad Hartlaub for his insight and guidance throughout this project, as well as, Chris Gillen for his assistance, the NSF (DMS - #0827208) and the Kenyon Summer Science Scholars Program